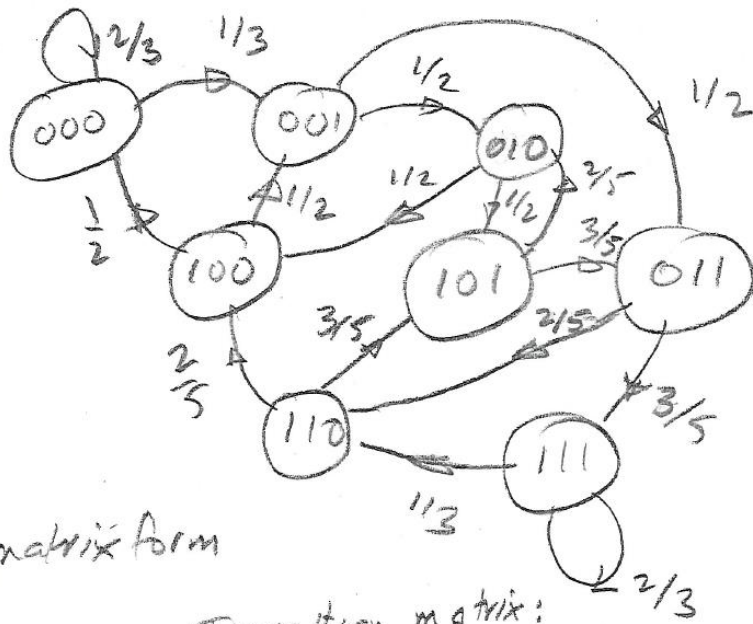


ECE 4260 Problem Set 8 Solutions

Problem 8.1

(a) I do not see any way to get by with less than 8 states:



(b) In matrix form

state

- 000 1
- 001 2
- 010 3
- 011 4
- 100 5
- 101 6
- 110 7
- 111 8

Transition matrix:

$$\begin{bmatrix}
 2/3 & 1/3 & 0 & 0 & 0 & 0 & 0 & 0 \\
 0 & 0 & 1/2 & 1/2 & 0 & 0 & 0 & 0 \\
 0 & 0 & 0 & 0 & 1/2 & 1/2 & 0 & 0 \\
 0 & 0 & 0 & 0 & 0 & 0 & 2/5 & 3/5 \\
 1/2 & 1/2 & 0 & 0 & 0 & 0 & 0 & 0 \\
 0 & 0 & 1/2 & 1/2 & 0 & 0 & 0 & 0 \\
 0 & 0 & 0 & 0 & 2/5 & 3/5 & 0 & 0 \\
 0 & 0 & 0 & 0 & 0 & 0 & 1/3 & 2/3
 \end{bmatrix}$$

= A

$$A^T \underline{P}_k = \underline{P}_{k+1}$$

FOR LIMITING STATE, $A^T \underline{P} = \underline{P}$ or $(A^T - I) \underline{P} = \underline{0}$

Idea: $P_1 = \frac{2}{3} P_1 + \frac{1}{2} P_5$

$$P_2 = \frac{1}{3} P_1 + \frac{1}{2} P_5$$

$$P_3 = \frac{1}{2} P_2 + \frac{2}{5} P_6$$

$$P_4 = \frac{1}{2} P_2 + \frac{3}{5} P_6$$

$$P_5 = \frac{1}{2} P_3 + \frac{2}{5} P_7$$

$$P_6 = \frac{1}{2} P_3 + \frac{3}{5} P_7$$

$$P_7 = \frac{2}{5} P_4 + \frac{1}{3} P_8$$

$$P_8 = \frac{3}{5} P_4 + \frac{2}{3} P_8$$

we also have

$$P_1 + P_2 + \dots + P_8 = 1$$

The eq's to the left are not enough.

$A^T - I$ is singular.

we replace any row with $(1, \dots, 1)$ and adjust the RHS.

If we eliminate the last eq & substitute $P_1 + P_8 = 1$, with a RHS of $[0, \dots, 0, 1]^T$,

we can solve the eq's. (Alternatively,

since A is sparse, we could solve by hand.)

Solution: $\underline{P}^T = [0.1429, 0.0952, 0.0952, 0.1190, 0.0952, 0.1190, 0.1190, 0.2143]$

$$P(\text{success}) = \left(\frac{1}{3}, \frac{1}{2}, \frac{1}{2}, \frac{1}{2}, \frac{3}{5}, \frac{3}{5}, \frac{3}{5}, \frac{2}{3} \right) \underline{P}$$

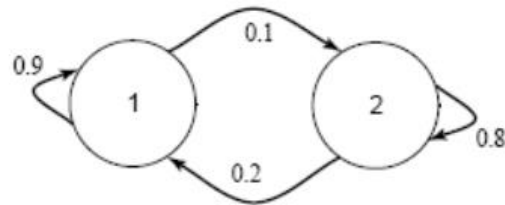
$$= 0.5452$$

Problem 8.2 (8.39 in Stark and Woods)

(a) With the two states $X = 1, 2$, we have state probability vector \mathbf{p} at time n (≥ 0) given as

$$\begin{aligned} \mathbf{p}[n] &= (P[X[n] = 1, P[X[n] = 2]) \\ &= (P[X[n-1] = 1]p_{11} + P[X[n-1] = 2]p_{21}, P[X[n-1] = 1]p_{12} + P[X[n-1] = 2]p_{22}) \\ &= \mathbf{p}[n-1]\mathbf{P}, \quad \text{with } \mathbf{P} \text{ the state-transition matrix,} \\ &= \mathbf{p}[n-2]\mathbf{P}^2, \\ &\quad \vdots \\ &= \mathbf{p}[0]\mathbf{P}^n. \end{aligned}$$

(b)



(c) Let p be the probability of the event {first transition to state 2 occurring at time n }. Then, given $X[0] = 1$, we have

$$\begin{aligned} p &= p_{11}^{n-1}p_{12} \\ &= (0.9)^{n-1}(0.1) \\ &= (0.1)(0.9)^{n-1}. \end{aligned}$$

Problem 8.3 (8.40 in Stark and Woods)

(a)

$$H(\omega) = \frac{1}{1 - re^{-j\omega}} \quad \text{and} \quad h[n] = r^n u[n],$$

so

$$\begin{aligned} S_{XX}(\omega) &= |H(\omega)|^2 S_{ZZ}(\omega) \\ &= \frac{1}{|1 - re^{-j\omega}|^2} \sigma_Z^2 \\ &= \frac{\sigma_Z^2}{1 + r^2 - 2r \cos \omega}. \end{aligned}$$

(b) We know $R_{XX}[m] = (h[m] * h^*[-m]) * \sigma_Z^2 \delta[m]$. Here, we have

$$\begin{aligned}
 h[n] * h^*[-n] &= \sum_{k=-\infty}^{+\infty} h[k] h^*[-(n-k)] \\
 &= \sum_{k=-\infty}^{+\infty} r^k u[k] r^{-(n-k)} u[k-n] \\
 &= r^{-n} \sum_{k=0}^{+\infty} r^{2k} u[k-n] \\
 &= \begin{cases} \frac{r^n}{1-r^2}, & n \geq 0, \\ \frac{r^{-n}}{1-r^2}, & n \leq 0, \end{cases} \\
 &= \frac{r^{|n|}}{1-r^2}, \quad \text{for all } n.
 \end{aligned}$$

Thus

$$\begin{aligned}
 R_{XX}[m] &= (h[m] * h^*[-m]) * \sigma_Z^2 \delta[m] \\
 &= \frac{r^{|m|}}{1-r^2} * \sigma_Z^2 \delta[m] \\
 &= \left(\frac{r^{|m|}}{1-r^2} * \delta[m] \right) \sigma_Z^2 \\
 &= \frac{r^{|m|}}{1-r^2} \sigma_Z^2.
 \end{aligned}$$

Problem 8.4 (8.42 in Stark and Woods)

(8.42 in S & W)

(a) FOR ANY INDEPENDENT INCREMENTS PROCESS:
 where $\Delta x(0) = 0$; $\Delta x(n) = \sum_{k=1}^n w_k(n)$; and $w_k(n)$ is
 uncorrelated noise with $\overline{w_k(n)} = \bar{w}$; $\overline{w_k^2(n)} = \bar{w}^2$;
 and $\sigma_{w_k(n)}^2 = \sigma_w^2$:

$$E(\Delta x(n)) = n, \bar{w} = \mu_{\Delta x(n)}$$

Here: $\bar{w} = \frac{1}{2} (s_1 - s_2)$

(b) $R_{\Delta x \Delta x}(n_1, n_2) = E(\Delta x(n_1) \Delta x(n_2))$

Special cases: let $n_1 = n_2$

$$R_{\Delta x \Delta x}(n_1, n_1) = \overline{\Delta x^2(n_1)} = E \left\{ \sum_{i=1}^{n_1} w(i) \sum_{j=1}^{n_1} w(j) \right\}$$

$$= \underbrace{\sum_{i=1}^{n_1} \sum_{j=1}^{n_1} \overline{w(i) w(j)}}_{n_1^2 \text{ terms}} = \underbrace{n_1 \bar{w}^2}_{\text{when } i=j} + \underbrace{n_1(n_1-1) \bar{w}^2}_{\text{when } i \neq j}$$

$$= n_1 (\bar{w}^2 + (n_1 - 1) \bar{w}^2)$$

Let $n_1 \neq n_2$ $n_2 > n_1$ w/o loss of generality.

$$E(\bar{\Delta}(n_1) \bar{\Delta}(n_2)) = E\left[\bar{\Delta}(n_1) \left(\bar{\Delta}(n_1) + \sum_{n=n_1+1}^{n_2} w(n)\right)\right]$$

$$= \overline{\Delta^2(n_1)} + \overline{\Delta(n_1)} \cdot (n_2 - n_1) \bar{w}$$

$$= n_1 \bar{w}^2 + (n_1)(n_1 - 1) \bar{w}^2 + n_1 \bar{w} (n_2 - n_1) \bar{w}$$

$$= n_1 \bar{w}^2 + [(n_1)(n_1 - 1) + n_1(n_2 - n_1)] \bar{w}^2$$

$$= n_1 (\bar{w}^2 [n_1 - 1 + n_2 - n_1] \bar{w}^2)$$

$$= n_1 (\bar{w}^2 + (n_2 - 1) \bar{w}^2)$$

$$\bar{w} = \frac{1}{2}(s_1 - s_2)$$

$$\bar{w}^2 = \frac{1}{2}(s_1^2 + s_2^2)$$